

ANOVA

$$(y_{ij} - Y_i) = (y_{ij} - \hat{Y}_i) + (\hat{Y}_i - Y_i) = (y_{ij} - \hat{\mu} - \hat{\alpha}_i) + (\hat{\mu} - \mu) + (\alpha_i - \hat{\alpha}_i)$$

$$\sum_i \sum_j (y_{ij} - Y_i)^2 = \sum_i \sum_j (y_{ij} - \hat{\mu} - \hat{\alpha}_i)^2 + (\hat{\mu} - \mu)^2 \sum_i p_i + \sum_i p_i (\hat{\alpha}_i - \alpha_i)^2 =$$

$$= \sum_i \sum_j (y_{ij} - y_{i.})^2 + (y_{..} - \mu)^2 \sum_i p_i + \sum_i p_i (y_{i.} - y_{..} - \alpha_i)^2$$

Fisher-Cochran-tétel $\sum_i p_i = \sum_i (p_i - 1) + 1 + (r - 1)$

mind $\chi^2 \sigma_e^2$

$$\sum_i p_i (y_{i.} - y_{..} - \alpha_i)^2 = \chi^2 \sigma_e^2$$

$$H_0 : \alpha_i = 0 \quad \longrightarrow \quad S_A = \sum_i p_i (y_{i.} - y_{..})^2 = \chi^2 \sigma_e^2$$

$$S_R = \sum_i \sum_j (y_{ij} - y_{i.})^2 = \chi^2 \sigma_e^2$$

$$s_A^2 = \frac{\sum_{i=1}^r p_i (y_{i.} - y_{..})^2}{r - 1} = \frac{\chi^2 \sigma_e^2}{V_A}$$

$$s_R^2 = \frac{\sum_i \sum_j (y_{ij} - y_{i.})^2}{\sum_i p_i - r} = \frac{\chi^2 \sigma_e^2}{V_R}$$

$$F_0 = \frac{s_A^2}{s_R^2}$$

ANOVA

Két faktor szerinti ANOVA

Az *A* faktor minden szintjét kombináljuk a *B* faktor minden szintjével, minden „cellában” azonos számú ismétlés (kiegyensúlyozott terv).

A terv szerkezete miatt a faktorok hatását egymásétól függetlenül vizsgálhatjuk.

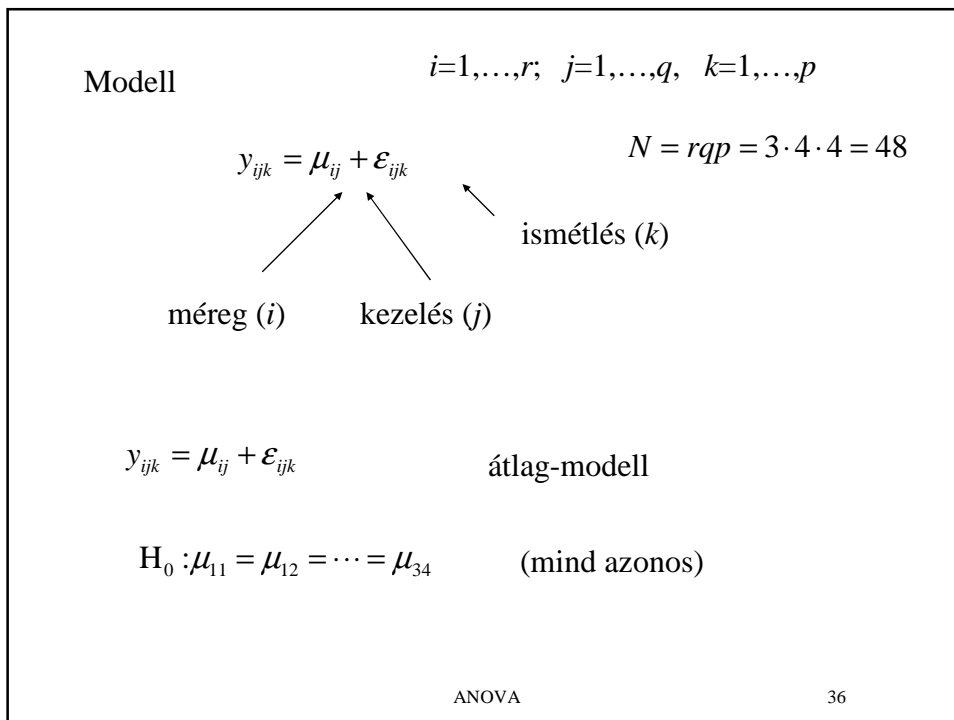
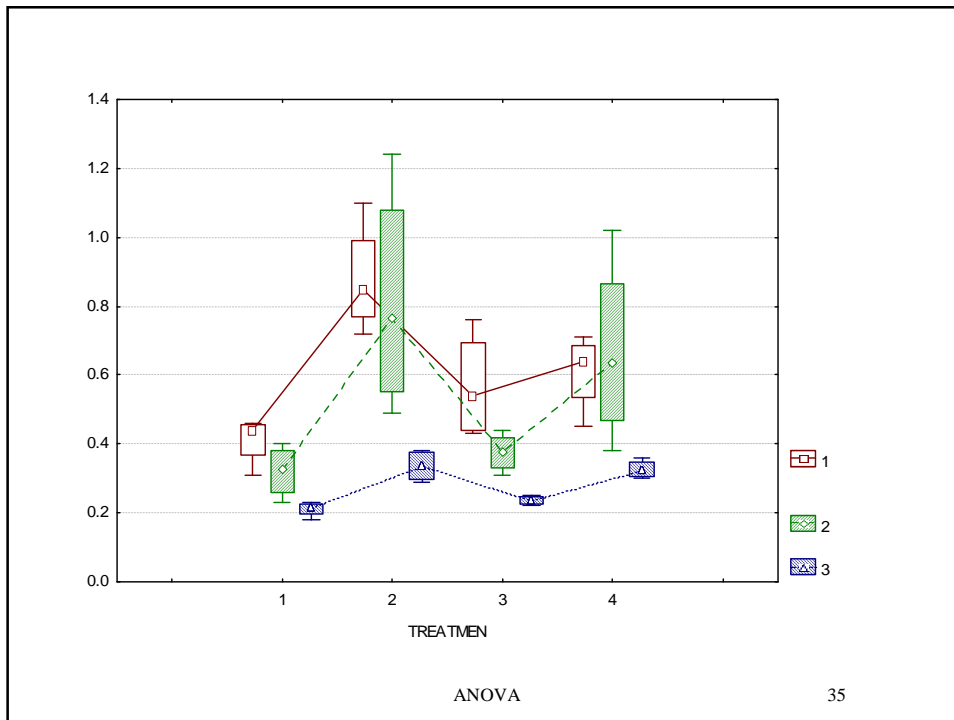
2. példa

(Box-Hunter-Hunter: Statistics for Experimenters, J. Wiley, 1978, p. 165)

poison.sta

		treatment			
		A	B	C	D
poison	I	0.310	0.820	0.430	0.450
		0.450	1.100	0.450	0.710
		0.460	0.880	0.630	0.660
		0.430	0.720	0.760	0.620
	II	0.360	0.920	0.440	0.560
		0.290	0.610	0.350	1.020
		0.400	0.490	0.310	0.710
		0.230	1.240	0.400	0.380
	III	0.220	0.300	0.230	0.300
		0.210	0.370	0.250	0.360
		0.180	0.380	0.240	0.310
		0.230	0.290	0.220	0.330

ANOVA



ANOVA

$$y_{ijk} = \mu + \alpha_i + \beta_j + \alpha\beta_{ij} + \varepsilon_{ijk}$$

↑

az i -edik méreg
hatása

↑

a j -edik kezelés
hatása

↑

kölcsönhatás

$y_{ijk} = \mu + \alpha_i + \beta_j + \alpha\beta_{ij} + \varepsilon_{ijk}$ hatás-modell

$H_0^A : \alpha_i = 0, i = 1, \dots, r$

$H_0^B : \beta_j = 0, j = 1, \dots, q$

$H_0^{AB} : \alpha\beta_{ij} = 0, i = 1, \dots, r; j = 1, \dots, q$

ANOVA
37

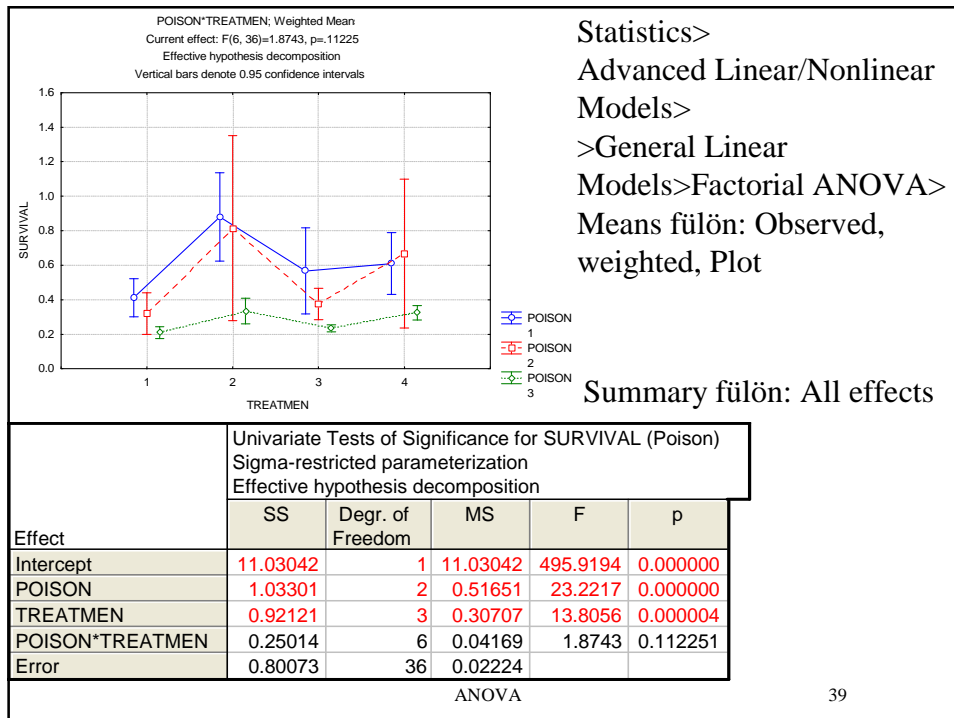
ANOVA-táblázat

Az eltérés forrása	eltérés-négyzetösszeg	szabadsági fok	szórásnégyzet	F
A hatása (sorok közötti)	$S_A = qp \sum_i (y_{i.} - y_{..})^2$	$r-1$	$s_A^2 = \frac{S_A}{r-1}$	S_A^2 / S_R^2
B hatása (oszlopok közötti)	$S_B = rp \sum_j (y_{.j} - y_{..})^2$	$q-1$	$s_B^2 = \frac{S_B}{q-1}$	S_B^2 / S_R^2
AB kölcsönhatás	$S_{AB} = p \sum_i \sum_j (y_{ij} - y_{i.} - y_{.j} + y_{..})^2$	$(r-1) \cdot (q-1)$	$s_{AB}^2 = \frac{S_{AB}}{(r-1)(q-1)}$	S_{AB}^2 / S_R^2
Maradék (csoportokon belüli)	$S_R = \sum_i \sum_j \sum_k (y_{ijk} - y_{ij.})^2$	$rq(p-1)$	$s_R^2 = \frac{S_R}{rq(p-1)}$	
Teljes	$S_0 = \sum_i \sum_j \sum_k (y_{ijk} - y_{...})^2$	$rqp-1$		

$3 \cdot 4 \cdot (4-1) = 36$

ANOVA
38

ANOVA



39

Homoszkedaszticitás

$$\sigma_e^2 = \text{konst} \quad ?$$

More results>Assumptions fülön: Homogeneity of variances

Tests of Homogeneity of Variances (Poison)					
Effect: POISON*TREATMEN					
	Hartley F-max	Cochran C	Bartlett Chi-Sqr.	df	p
SURVIVAL	678.6000	0.423741	45.13689	11	0.000005

Levene's Test for Homogeneity of Variances (Poison)				
Effect: POISON*TREATMEN				
Degrees of freedom for all F's: 11, 36				
	MS Effect	MS Error	F	p
SURVIVAL	0.024601	0.005069	4.853537	0.000144

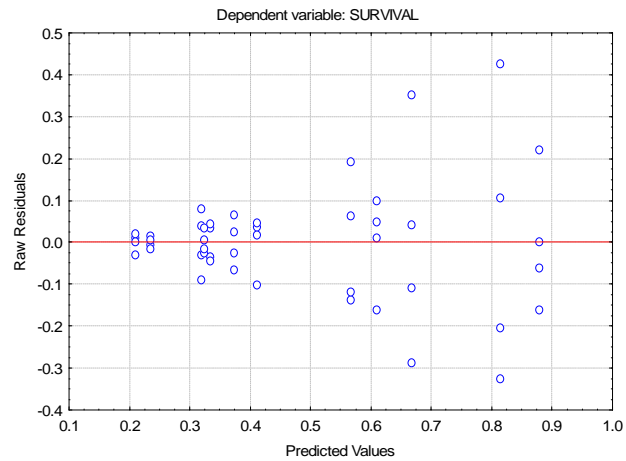
ANOVA

40

ANOVA

A reziduumok vizsgálata

Residuals1 fülön: Pred. & resid.



ANOVA

41

Box-Cox transzformáció

$$\sigma_y \sim \bar{y}^\alpha \quad \text{Var}(y^{tr}) = \left(\frac{dy^{tr}}{dy}\right)^2 \sigma_y^2 = \left(\frac{dy^{tr}}{dy}\right)^2 y^{2\alpha}$$

$$\text{Var}(y^{tr}) = \text{konst} \quad \text{ha} \quad \frac{dy^{tr}}{dy} y^\alpha = \text{konst}$$

$$dy^{tr} = ky^{-\alpha} dy$$

$$y^{tr} = \int y^{-\alpha} dy = \begin{cases} y^{1-\alpha} & \text{ha } \alpha \neq 1 \\ \ln y & \text{ha } \alpha = 1 \end{cases}$$

ANOVA

42

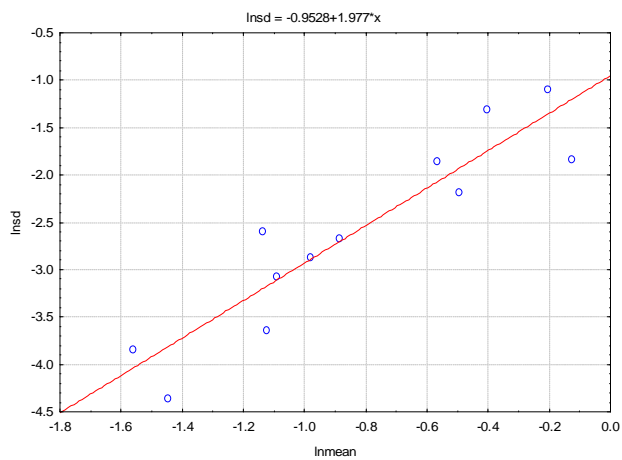
ANOVA

$$y^{\alpha} = \int y^{-\alpha} dy = \begin{cases} y^{1-\alpha} & \text{ha } \alpha \neq 1 \\ \ln y & \text{ha } \alpha = 1 \end{cases}$$

α	$\lambda=1-\alpha$	transzformáció
2	-1	$1/y$
1.5	-0.5	$1/\sqrt{y}$
1	0	$\ln y$
0.5	0.5	\sqrt{y}
0	1	(nincs transzformáció)



$\sigma_y \sim \bar{y}^\alpha$ $\ln \sigma_y = k + \alpha \ln \bar{y}$ egyenest kell illeszteni



$\alpha \approx 2$

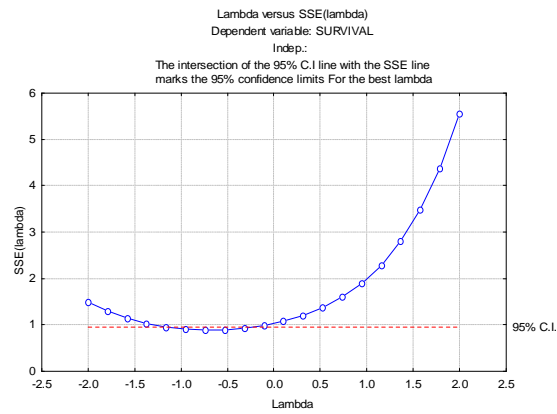
$\sigma_y \sim \bar{y}^2$

ANOVA

Box-Cox transzformáció

$$y^{tr} = y^\lambda$$

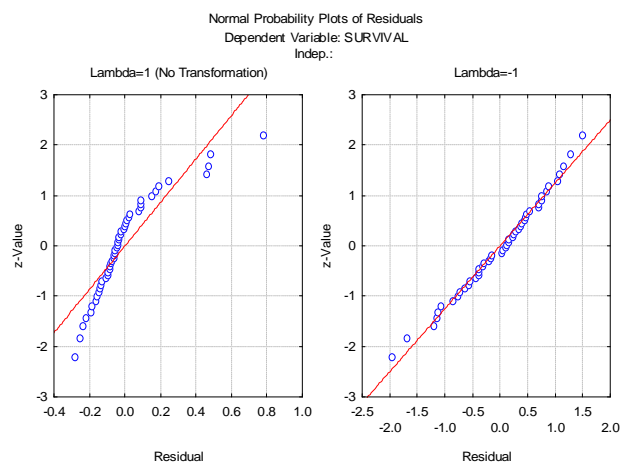
File>Open: (Program Files>StatSoft>Statistica8>Examples>Macros>
>Analysis Examples>BoxCox)



$$\lambda \approx -1$$

ANOVA

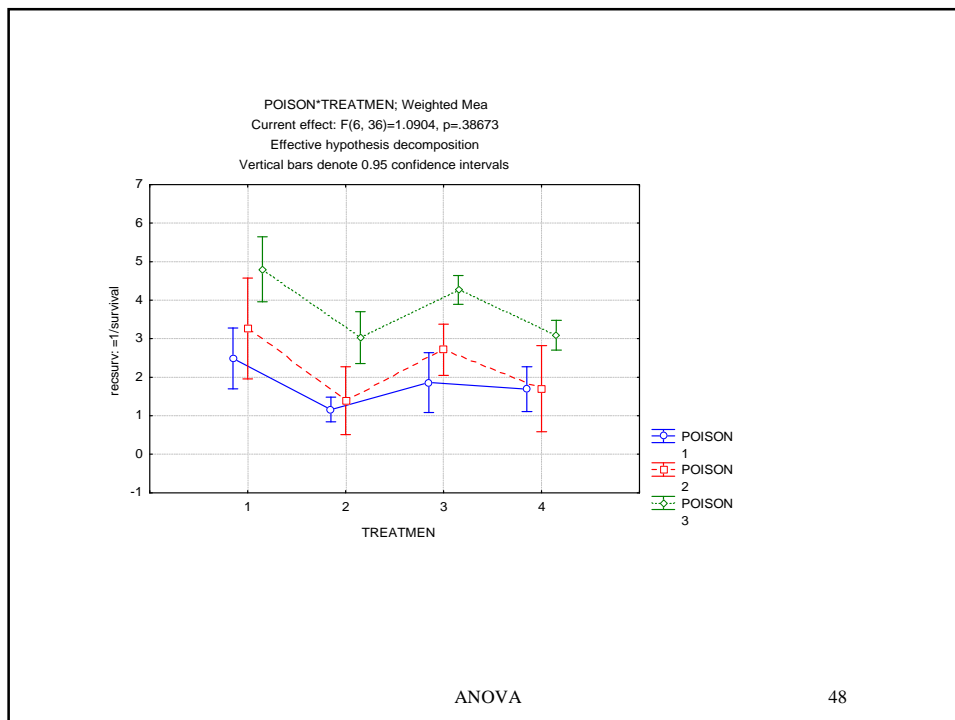
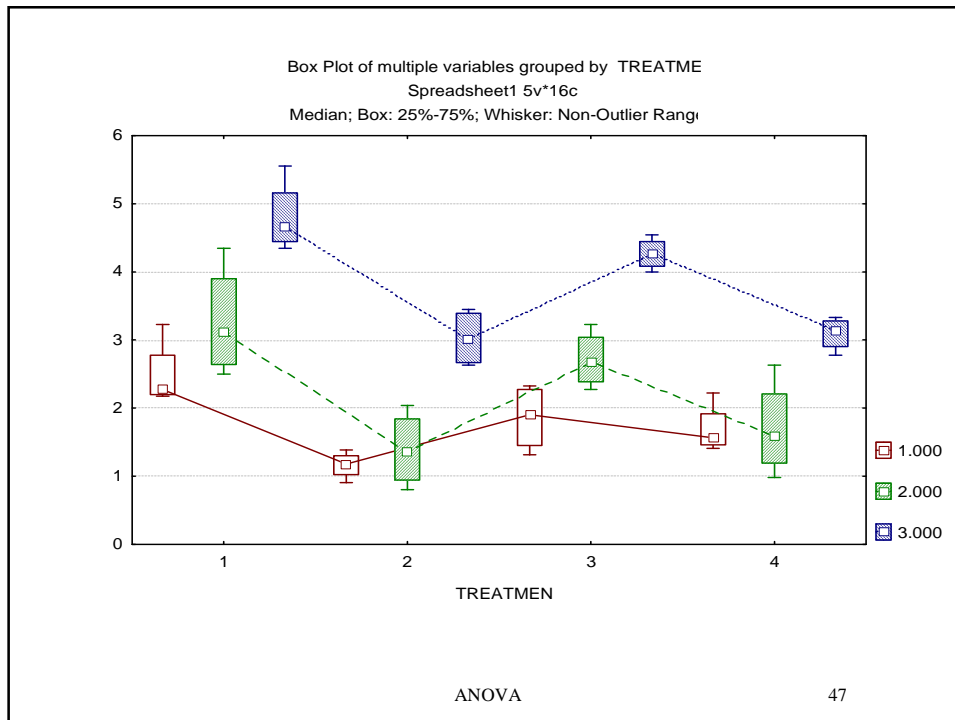
45



ANOVA

46

ANOVA



ANOVA

Univariate Tests of Significance for Recsurv (Poison)					
Sigma-restricted parameterization					
Effective hypothesis decomposition					
Effect	SS	Degr. of Freedom	MS	F	p
Intercept	330.0892	1	330.0892	1374.881	0.000000
POISON	34.8771	2	17.4386	72.635	0.000000
TREATMEN	20.4143	3	6.8048	28.343	0.000000
POISON*TREATMEN	1.5708	6	0.2618	1.090	0.386733
Error	8.6431	36	0.2401		

A hatások még kifejezettebbek (F értékei nagyobbak), a kölcsönhatáshoz tartozó p 0.112 helyett 0.387 lesz.

A reziduumok vizsgálata

